# Introduction to Confidence Intervals 

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1. General idea, the CI based on Z <br> 2. The t-Student distribution <br> 3. t-Student's Cl for the population mean <br> 4. Two (independent) samples: t-Student's Cl for the mean difference
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## Cl when population standard deviation $\sigma$ is known

- If either $X_{1}, \ldots, X_{n}$ are normal distributed, or $n$ is so large that the central limit theorem starts to work, then

$$
\frac{\bar{X}-\mu}{\sigma / \sqrt{n}} \sim N(0,1)
$$

- This means that we can produce a $95 \%$ confidence interval as follows:

$$
95 \% \mathrm{Cl}=\left(\bar{X}-1.96 \times \frac{\sigma}{\sqrt{n}}, \bar{X}+1.96 \times \frac{\sigma}{\sqrt{n}}\right)
$$

where 1.96 is the two-sided $\mathbf{5 \%}$ point of the standard normal distribution.

- Small $\sigma$ or large $n$ gives more narrow intervals and means that $\bar{X}$ is more likely to be similar to the true mean.


## Unknown $\sigma$ and the Student t -distribution

- The previous situation with known $\sigma$ is rare in practice,
- We will use the previous strategy, but replace $\sigma$ with the empirical standard deviation

$$
s=\sqrt{\frac{1}{n-1} \sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}}
$$

- The added uncertainty means that $\frac{\bar{X}-\mu}{s / \sqrt{n}}$ is not normal distributed anymore, but distributed according to the so-called Student t-distribution with n-1 degrees of freedom, written

$$
\frac{\bar{X}-\mu}{s / \sqrt{n}} \sim t(n-1)
$$



Figur 8.3 Standardnormalfordelingen er tegnet inn sammen med Studentfordelingen med 3 frihetsgrader. En ser at den siste fordelingen er mer spredt ut enn den første

- The higher degree of freedom the closer the stundent t is to the standard normal distribution $\mathrm{N}(0,1)$


Figur 8.4 Det er gjort 1000 utvalg, hvert på fire tall, fra et sett med data over mannlig kroppshøyde. Størrelsen $t$ er beregnet i hvert av utvalgene, og histogrammet viser fordelingen av $t$-verdiene. Som sammenlikning vises den standardiserte normalfordelingen (prikket kurve) og Studentfordelingen med 3 frihetsgrader (heltrukket kurve)

## $95 \%$ confidence interval when $\sigma$ is unknown

- Can use previous strategy, but we cannot use the percentage point 1.96 anymore, since $\frac{\bar{X}-\mu}{\sigma / \sqrt{n}}$ is not normally distributed
- We must use the percentage point for the Student t -distribution with n -1 degrees of freedom instead, $t^{\prime}$
- We still need that either $X_{1}, \ldots, X_{n}$ are normally distributed, or $n$ is so large that the central limit theorem ensures that $\bar{X}$ is normally distributed
- A 95\%-confidence interval is then given by:

$$
\mathrm{Cl}=\left(\bar{X}-t^{\prime} \times \frac{s}{\sqrt{n}}, \bar{X}+t^{\prime} \times \frac{s}{\sqrt{n}}\right)
$$

where s denotes the empirical standard deviation, and $\bar{X}$ denotes the sample mean.

## Example 6.3 in Kirkwood \& Sterne

The numbers of hours of relief obtained by six arthritic patients after receiving a new drug are recorded.

| Patient no. | 1 | 2 | 3 | 4 | 5 | 6 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Number of hours | 2.2 | 2.4 | 4.9 | 2.5 | 3.7 | 4.3 |

The sample mean is $\bar{X}=3.3$ hours, the empirical standard deviation is $s=1.13$ hours and the estimated standard error of the sample mean equals $s / \sqrt{n}=0.46$ hours. The number of degrees of freedom is $(n-1)=5$

The 95\% confidence interval (in hours) for the average number of hours of relief for arthritic patients in general is

$$
(3.3-2.57 \times 0.46,3.3+2.57 \times 0.46)=(2.1,4.5),
$$

where 2.57 is the two-sided $5 \%$ point of the $t$ distribution with 5 degrees of freedom

## We can simply use the normal distribution when n is large

- When $n$ is large ( 150 or larger), then $t^{\prime}$ is almost the same as 1.96 for practical purposes,
- Reflects that the Student t-distribution is almost identical to the standard normal distribution for large degrees of freedom,
- The $95 \%$ confidence interval for the population mean is then given by

$$
95 \% \mathrm{Cl}=\left(\bar{X}-1.96 \times \frac{s}{\sqrt{n}}, \bar{X}+1.96 \times \frac{s}{\sqrt{n}}\right) .
$$

## Example 6.1 in Kirkwood \& Sterne

We want to estimate the amount of insecticide that would be required to spray all the 10000 houses in a rural area as part of a malaria control programme. A random sample of 100 houses is chosen and the sprayable surface of each of these is measured. The mean sprayable surface area for these 100 houses is $\bar{X}=24.2$ $\mathrm{m}^{2}$, and the estimated standard deviation is $s=5.9 \mathrm{~m}^{2}$.
The estimated standard error of the sample mean is $s / \sqrt{n}=5.9 / \sqrt{100}=0.6 \mathrm{~m}^{2}$.

The $95 \%$ confidence interval is:

$$
(24.2-1.96 \times 0.6,24.2+1.96 \times 0.6)=(23.0,25.4)
$$

The upper 95\% confidence limit is used in budgeting for the amount of insecticide required per house. One litre of insecticide is sufficient to spray $50 \mathrm{~m}^{2}$ and so the amount (in litres) budgeted for is:

$$
10000 \times \frac{25.4}{50}=5080
$$

## Small sample sizes

- The central limit theorem says that $\bar{X}$ is normally distributed, even if the individual observations are not
- For smaller samples, we need that the individual samples are normally distributed
- This can be easily checked with a normality plot in $\mathbf{R}$
- When the distribution in the population is markedly non-normal, it may be desirable to
- use a transformation on the scale on which the variable $X$ is measured, or
- calculate a non-parametric confidence interval, or
- use bootstrap methods

More on this in day 1 of week 2 !

## Confidence interval vs. reference range

- If the population distribution is approximately normal, the $95 \%$ reference range is given by

$$
95 \% \text { reference range }=(\mu-1.96 \times \sigma, \mu+1.96 \times \sigma)
$$

where $\mu$ is the population mean and $\sigma$ is the population standard deviation

- There is a clear distinction between the Cl and the reference range:
- the reference range describes the variability between individual observations in the population
- the confidence interval is a range of plausible values for the population mean, given the sample mean and its standard error Since the sample size $n>1$, the confidence interval will always be narrower than the reference range.


## What if we have more than one Sample?

Two Independent Samples
2 groups: 1 measure for each individual, each which corresponds to a group (for example sick/healthy people)

| Group 1 |  | Group 2 |  |
| :--- | :--- | :--- | :--- |
| Ind. | Measure | Ind. | Measure |
| 1 | $X_{11}$ | 1 | $X_{12}$ |
| 2 | $X_{21}$ | 2 | $X_{22}$ |
| $\ldots$ |  | $\ldots$ |  |
|  |  | 14 | $X_{142}$ |
| 15 | $X_{151}$ |  |  |

## The mean difference of two independent samples

- We want to compare the mean outcomes in two separate exposure (or treatment) groups: group 0 and group 1
- In clinical trials, these correspond to the treatment and control groups,
- We will then build a two-sample confidence interval,
- Notation:
$n_{i}$ is number of individuals in group $i$,
- $X_{1, i}, \ldots, X_{n_{i}, i}$ observations in group $n_{i}$,
- $\bar{X}_{i}$ average in group $i$,
- $\mu_{i}$ mean in group $i$,
- $\sigma_{i}$ standard deviation in group $i$,
- $s_{i}$ empirical standard deviation in group $i$.


## Assumptions and the Student t-distribution

- Independent individuals,
- Normal distributed averages, i.e. either
- Large enough samples such that averages become normal distributed, or
- Normal distributed observations,
- Equality of the two population standard deviations, $\sigma_{1}$ and $\sigma_{0}$

This means that

$$
\begin{equation*}
t=\frac{\bar{X}_{1}-\bar{X}_{0}}{s \sqrt{\left(1 / n_{1}+1 / n_{0}\right)}} \tag{1}
\end{equation*}
$$

is t -distributed with $n_{1}+n_{0}-2$ degrees of freedom, where

$$
\begin{equation*}
s=\sqrt{\left[\frac{\left(n_{1}-1\right) s_{1}^{2}+\left(n_{0}-1\right) s_{0}^{2}}{\left(n_{1}+n_{0}-2\right)}\right]} \tag{2}
\end{equation*}
$$

## Confidence interval for the mean difference $\mu_{1}-\mu_{0}$

- The confidence interval gives a range of likely values for the difference in population means, $\mu_{1}-\mu_{0}$, based on the difference in sample means, $\bar{X}_{1}-\bar{X}_{0}$ :

$$
\mathrm{Cl}=\left(\bar{X}_{1}-\bar{X}_{0}\right) \pm t^{\prime} \times s \sqrt{1 / n_{1}+1 / n_{0}}
$$

where the common estimate, $s$, of the population standard deviation is given by (also at the previous slide):

$$
s=\sqrt{\left[\frac{\left(n_{1}-1\right) s_{1}^{2}+\left(n_{0}-1\right) s_{0}^{2}}{\left(n_{1}+n_{0}-2\right)}\right]}
$$

and $t^{\prime}$ is the appropriate percentage point of the $t$ distribution with $\left(n_{1}+n_{0}-2\right)$ degrees of freedom

## Example: 7.2 in Kirkwood \& Sterne

We consider the birth weights (in kg ) of children born to 14 heavy smokers (group 1) and to 15 non-smokers (group 0), sampled from live births at a large teaching hospital

| Heavy smokers <br> (group 1) | Non-smokers <br> (group 0) |
| :---: | :---: |
| 3.18 | 3.99 |
| 2.74 | 3.89 |
| 2.90 | 3.60 |
| 3.27 | 3.73 |
| 3.65 | 3.31 |
| 3.42 | 3.70 |
| 3.23 | 4.08 |
| 2.86 | 3.61 |
| 3.60 | 3.83 |
| 3.65 | 3.41 |
| 3.69 | 4.13 |
| 3.53 | 3.36 |
| 2.38 | 3.54 |
| 2.34 | 3.51 |
|  | 2.71 |
| $\bar{X}_{1}=3.1743$ | $\bar{X}_{0}=3.6267$ |
| $s_{1}=0.4631$ | $s_{0}=0.3584$ |
| $n_{1}=14$ | $n_{0}=15$ |

The difference between the means is given by

$$
\bar{X}_{1}-\bar{X}_{0}=3.1743-3.6267=-0.4524
$$

and the standard deviation is given by

$$
s=\sqrt{\frac{13 \times 0.4631^{2}+14 \times 0.3584^{2}}{14+15-2}}=0.4121
$$

with $(14+15-2)=27$ degrees of freedom. The standard error of the difference is given by

$$
\widehat{\text { s.e. }}=0.4121 \times \sqrt{(1 / 14+1 / 15)}=0.1531
$$

The 95\% confidence interval for the difference between the mean birth weight is given by

$$
\begin{aligned}
(-0.4524-2.05 \times 0.1531,-0.4524+2.05 & \times 0.1531) \\
& =(-0.77,-0.14)
\end{aligned}
$$

where 2.05 is the $5 \%$ point of the $t$ distribution with 27 degrees of freedom

